

ZERO-INFLATED POISSON TRANSMUTED WEIGHTED EXPONENTIAL DISTRIBUTION: PROPERTIES AND APPLICATIONS

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ABSTRACT. *Count observations with high frequencies of zero counts abound in diverse fields. In actuary science, for example, insurance claims are often underreported, leading to a higher frequency of zero counts. This invariably reduces the mean and leads to over-dispersion. Different techniques to model such occurrences exist. This study uses the cubic rank transmutation map to compound the weighted exponential distribution and obtain a new count distribution in the mixed Poisson paradigm. The zero-inflated form of the proposition is obtained along with some mathematical properties. Simulated skewness, kurtosis, and dispersion index for the new distributions reveal they are suitable for model dispersed observation with positive skewness. Five datasets with high frequencies of zero counts are used to assess the performance of the new distribution along with some popular count distributions by using the maximum likelihood for parameter estimation. Results show that the natural form of the new proposition performs creditably better than its zero-inflated forms, even when there is a higher proportion of zero counts. The classical negative binomial distribution is also observed to outperform its zero-inflated form in most cases. In contrast, the zero-inflated Poisson distribution better fits the datasets than the classical Poisson distribution.*

KEYWORDS: Mixed Poisson distribution, mixing distribution, rank transmutation map, weighted exponential distribution, excess zero counts.

INTRODUCTION

Results obtained from count data modelling may be misleading if there are too many or too few zero counts. Zero-deflation arises when the zero frequency in a dataset is lower than the expected frequency, while zero inflation occurs when there are too many zero counts. The former is rare, while the latter is more observed (Conceição *et al.*, 2017). The Poisson distribution is often considered for modelling count data (Tajuddin & Ismail, 2022; Wagh & Kamalja, 2017). The distribution assumes equality of mean and variance for observations (Ong *et al.*, 2021). Observations with a relatively higher frequency of zero counts are usually dispersed (Sellers & Raim, 2016). Since most count data are dispersed (Adetunji &

Sabri, 2021), there is model misspecification when the classical Poisson distribution is assumed for dispersed observation (Asamoah, 2016). In order to overcome challenges that characterize the Poisson distribution, several methods that can handle dispersed count observations with excess zero have been proposed (Das *et al.*, 2018; Ong *et al.*, 2021). One of the techniques often utilized for modelling such observation is the mixed Poisson, first proposed in the early 20th century (Greenwood & Yule, 1920) when the gamma distribution is assumed for the parameter of the classical Poisson distribution resulting in the negative binomial distribution. The process of obtaining a mixed Poisson distribution involves assuming a continuous distribution (called the mixing distribution) with positive supports for the parameter of the Poisson distribution.

Several choices of mixing distributions have been proposed to improve the flexibility and general applicability of the mixed Poisson paradigm. A detailed survey on the choice of the mixing distribution is provided in Ong *et al.* (2021), while Karlis & Xekalaki (2005) gave several properties of the mixed Poisson distribution.

In this study, the cubic rank transmutation map (Rahman *et al.*, 2019) is used to extend the weighted exponential distribution (Gupta & Kundu, 2009) and obtain a new mixing distribution assumed for the parameter of the Poisson distribution. The mixing distribution is used to obtain a new mixed Poisson distribution and its zero-inflated form.

Zero-inflated distribution is often used to model observations with several zeros by assigning an extra probability to zero occurrences (Lambert, 1992). The distribution is applied when the frequency of excess zeros is assumed to have come from two processes; the first, where zero counts are obtained by chance like the ones, twos, etc., and the second are obtained when some data are constrained to be zeros (Lambert, 1992; Shahmandi *et al.*, 2020). The claim frequency in actuary science is an example of such modelling since observed zeros could have been from two scenarios. A policyholder may have no claim (no case of an accident) reflecting a true zero. Also, there are situations when the policyholder may be involved in a minor accident and hence, may have no urge to report such for a claim, mainly due to the usually cumbersome processes and procedures involved in getting the claim.

MATERIALS AND METHOD

Weighted Exponential Distribution

The weighted exponential distribution (Gupta & Kundu, 2009) with the CDF defined in equation (1) as:

$$G(t) = 1 - \frac{1}{a} e^{-\theta t} (a + 1 - e^{-a\theta t}) \quad (1)$$

Gupta and Kundu (2009) showed that the shape of equation (1) is identical to that of other two-parameter continuous distributions like generalized exponential, Weibull, and gamma, hence can be used as their alternative. The 3-parameter form of the distribution was proposed by Altun (2019). Zamani *et al.* (2014) used the mixing distribution to obtain a new mixed Poisson distribution with applications in claim frequency.

Cubic Rank Transmutation Map

Several extensions of different baseline distributions have been proposed to improve flexibility and general applicability. Among many recently proposed techniques are: the Quadratic Transmutation (QT) map (Shaw & Buckley, 2007); and Cubic Rank Transmutations (CRT) map (Al-kadim, 2018; Aslam *et al.*, 2018; Rahman *et al.*, 2019). Given a baseline distribution with distribution function $G(t)$, the cubic rank transmutation map (Rahman *et al.*, 2019) has the CDF given in equation (2) as:

$$F(t) = (1 - p)G(t) + 3p(G(t))^2 - 2p(G(t))^3 \tag{2}$$

Inserting equation (1) into equation (2) gives the distribution function for the Cubic Transmuted Weighted Exponential Distribution (CTWED) in equation (3) as:

$$F(t) = \left(\frac{(a+1)e^{-\theta t} - e^{-(1+a)\theta t} - a}{a^3} \right) (pae^{-(1+a)\theta t} - 4p(a+1)e^{-(2+a)\theta t} + 2pe^{-2(1+a)\theta t} - pa(a+1)e^{-\theta t} + 2p(a+1)^2e^{-2\theta t} - a^2) \tag{3}$$

The corresponding PDF is obtained in equation (4) as:

$$f(t) = \frac{\theta(a+1)}{a^3} [-6p(2a+3)e^{-(2a+3)\theta t} + 6p(a+1)(a+3)e^{-(a+3)\theta t} + a^2(p-1)e^{-(a+1)\theta t} - 6pa(a+2)e^{-(a+2)\theta t} + 6pe^{-3(a+1)\theta t} + 6pae^{-2(a+1)\theta t} - 6p(a+1)^2e^{-3\theta t} + 6pa(a+1)e^{-2\theta t} - a^2(p-1)e^{-\theta t}] \tag{4}$$

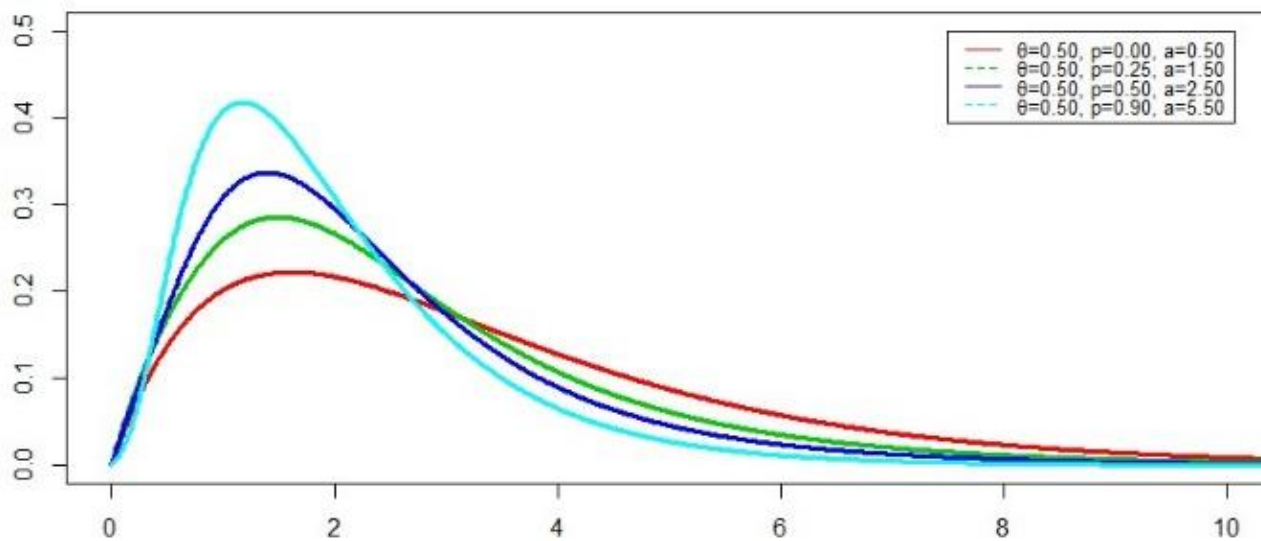


Figure 1. Shapes of the PDF of the CTWED

Figure 1 shows that the PDF of the CTWED is positively skewed and unimodal.

Moment and Moment Generating Function

Proposition 1: If a random variable T has a CTWED, the r^{th} moment is obtained in equation (5) as:

$$E(t^r) = \left(\frac{(a+1)r!}{a^3\theta^r} \right) \left(\frac{-6p}{(2a+3)^r} + \frac{6p(a+1)}{(a+3)^r} + \frac{a^2(p-1)}{(a+1)^{r+1}} - \frac{6pa}{(a+2)^r} + \frac{2p}{3^r(a+1)^{r+1}} + \frac{3pa}{2^r(a+1)^{r+1}} - \frac{2p(a+1)^2}{3^r} + \frac{3pa(a+1)}{2^r} - a^2(p-1) \right) \tag{5}$$

Proof

$$\begin{aligned} E(t^r) &= \int_0^\infty t^r f(t) dt \\ &= \int_0^\infty t^r \frac{\theta(a+1)}{a^3} [-6p(2a+3)e^{-(2a+3)\theta t} + 6p(a+1)(a+3)e^{-(a+3)\theta t} + a^2(p-1)e^{-(a+1)\theta t} \\ &\quad - 6pa(a+2)e^{-(a+2)\theta t} + 6pe^{-3(a+1)\theta t} + 6pae^{-2(a+1)\theta t} - 6p(a+1)^2e^{-3\theta t} \\ &\quad + 6pa(a+1)e^{-2\theta t} - a^2(p-1)e^{-\theta t}] dt \\ &= \frac{\theta(a+1)}{a^3} [-6p(2a+3) \int_0^\infty t^r e^{-(2a+3)\theta t} dt + 6p(a+1)(a+3) \int_0^\infty t^r e^{-(a+3)\theta t} dt + a^2(p-1) \int_0^\infty t^r e^{-(a+1)\theta t} dt \\ &\quad - 6pa(a+2) \int_0^\infty t^r e^{-(a+2)\theta t} dt + 6p \int_0^\infty t^r e^{-3(a+1)\theta t} dt + 6pa \int_0^\infty t^r e^{-2(a+1)\theta t} dt - 6p(a+1)^2 \int_0^\infty t^r e^{-3\theta t} dt \\ &\quad + 6pa(a+1) \int_0^\infty t^r e^{-2\theta t} dt - a^2(p-1) \int_0^\infty t^r e^{-\theta t} dt] \\ &= \left(\frac{\theta(a+1)r!}{a^3\theta^{r+1}} \right) \left(\frac{-6p(2a+3)}{(2a+3)^{r+1}} + \frac{6p(a+1)(a+3)}{(a+3)^{r+1}} + \frac{a^2(p-1)}{(a+1)^{r+1}} - \frac{6pa(a+2)}{(a+2)^{r+1}} + \frac{6p}{3^{r+1}(a+1)^{r+1}} + \frac{6pa}{2^{r+1}(a+1)^{r+1}} - \frac{6p(a+1)^2}{3^{r+1}} + \frac{6pa(a+1)}{2^{r+1}} - a^2(p-1) \right) \\ &= \left(\frac{(a+1)r!}{a^3\theta^r} \right) \left(\frac{-6p}{(2a+3)^r} + \frac{6p(a+1)}{(a+3)^r} + \frac{a^2(p-1)}{(a+1)^{r+1}} - \frac{6pa}{(a+2)^r} + \frac{2p}{3^r(a+1)^{r+1}} + \frac{3pa}{2^r(a+1)^{r+1}} - \frac{2p(a+1)^2}{3^r} + \frac{3pa(a+1)}{2^r} - a^2(p-1) \right) \end{aligned}$$

Proposition 2: If a random variable T has a CTWED, the MGF is obtained in equation (6) as:

$$E(e^{zt}) = \frac{\theta(a+1)}{a^3} \left(\frac{-6p(2a+3)}{2a\theta+3\theta-z} + \frac{6p(a+1)(a+3)}{a\theta+3\theta-z} + \frac{a^2(p-1)}{a\theta+\theta-z} - \frac{6pa(a+2)}{a\theta+2\theta-z} + \frac{6p}{3a\theta+3\theta-z} + \frac{6pa}{2a\theta+2\theta-z} - \frac{6p(a+1)^2}{3\theta-z} + \frac{6pa(a+1)}{2\theta-z} - \frac{a^2(p-1)}{\theta-z} \right) \tag{6}$$

Proof

$$\begin{aligned} E(e^{zt}) &= \int_0^\infty e^{zt} f(t) dt \\ &= \int_0^\infty e^{zt} \frac{\theta(a+1)}{a} [-6p(2a+3)e^{-(2a+3)\theta t} + 6p(a+1)(a+3)e^{-(a+3)\theta t} + a^2(p-1)e^{-(a+1)\theta t} - 6pa(a+2)e^{-(a+2)\theta t} + 6pe^{-3(a+1)\theta t} + 6pae^{-2(a+1)\theta t} - 6p(a+1)^2e^{-3\theta t} + 6pa(a+1)e^{-2\theta t} - a^2(p-1)e^{-\theta t}] dt \end{aligned}$$

$$\begin{aligned}
&= \frac{\theta(a+1)}{a^3} \int_0^{\infty} \left(-6p(2a+3)e^{-(2a\theta+3\theta-z)t} + 6p(a+1)(a+3)e^{-(a\theta+3\theta-z)t} \right. \\
&\quad \left. + a^2(p-1)e^{-(a\theta+\theta-z)t} - 6pa(a+2)e^{-(a\theta+2\theta-z)t} + 6pe^{-(3a\theta+3\theta-z)t} \right. \\
&\quad \left. + 6pae^{-(2a\theta+2\theta-z)t} - 6p(a+1)^2 e^{-(3\theta-z)t} + 6pa(a+1)e^{-(2\theta-z)t} \right. \\
&\quad \left. - a^2(p-1)e^{-(\theta-z)t} \right) dt \\
&= \frac{\theta(a+1)}{a^3} \left(\frac{-6p(2a+3)}{2a\theta+3\theta-z} + \frac{6p(a+1)(a+3)}{a\theta+3\theta-z} + \frac{a^2(p-1)}{a\theta+\theta-z} - \frac{6pa(a+2)}{a\theta+2\theta-z} + \frac{6p}{3a\theta+3\theta-z} \right. \\
&\quad \left. + \frac{6pa}{2a\theta+2\theta-z} - \frac{6p(a+1)^2}{3\theta-z} + \frac{6pa(a+1)}{2\theta-z} - \frac{a^2(p-1)}{\theta-z} \right)
\end{aligned}$$

Mixed Poisson CTWED

Proposition 3: Given that a random variable $X \sim \text{Poisson}(T)$ and $T \sim \text{CTWED}(\theta, p, a)$. A discrete random variable X has a mixed Poisson-CTWED (PCTWED) if its PMF is defined in equation (7) as:

$$\begin{aligned}
P_x &= \frac{\theta(a+1)}{a^3} \left(\frac{6p(a+1)(a+3)}{(1+a\theta+3\theta)^{x+1}} - \frac{6p(2a+3)}{(1+2a\theta+3\theta)^{x+1}} + \frac{a^2(p-1)}{(1+a\theta+\theta)^{a+1}} - \frac{6pa(a+2)}{(1+a\theta+2\theta)^{x+1}} + \frac{6p}{(1+3a\theta+3\theta)^{x+1}} + \right. \\
&\quad \left. \frac{6pa}{(1+2a\theta+2\theta)^{x+1}} - \frac{6p(a+1)^2}{(1+3\theta)^{x+1}} + \frac{6pa(a+1)}{(1+2\theta)^{x+1}} - \frac{a^2(p-1)}{(1+\theta)^{x+1}} \right) \quad (7)
\end{aligned}$$

Proof

$$\begin{aligned}
P_x &= \int_0^{\infty} \frac{t^x e^{-t}}{x!} \frac{\theta(a+1)}{a^3} \left[-6p(2a+3)e^{-(2a+3)\theta t} + 6p(a+1)(a+3)e^{-(a+3)\theta t} \right. \\
&\quad \left. + a^2(p-1)e^{-(a+1)\theta t} - 6pa(a+2)e^{-(a+2)\theta t} + 6pe^{-3(a+1)\theta t} + 6pae^{-2(a+1)\theta t} \right. \\
&\quad \left. - 6p(a+1)^2 e^{-3\theta t} + 6pa(a+1)e^{-2\theta t} - a^2(p-1)e^{-\theta t} \right] dt \\
&= \frac{\theta(a+1)}{a^3 x!} \int_0^{\infty} t^x \left(6p(a+1)(a+3)e^{-(1+a\theta+3\theta)t} - 6p(2a+3)e^{-(1+2a\theta+3\theta)t} \right. \\
&\quad \left. + a^2(p-1)e^{-(1+a\theta+\theta)t} - 6pa(a+2)e^{-(1+a\theta+2\theta)t} + 6pe^{-(1+3a\theta+3\theta)t} \right. \\
&\quad \left. + 6pae^{-(1+2a\theta+2\theta)t} - 6p(a+1)^2 e^{-(1+3\theta)t} + 6pa(a+1)e^{-(1+2\theta)t} \right. \\
&\quad \left. - a^2(p-1)e^{-(1+\theta)t} \right) dt \\
&= \frac{\theta(a+1)}{a^3} \left(\frac{6p(a+1)(a+3)}{(1+a\theta+3\theta)^{x+1}} - \frac{6p(2a+3)}{(1+2a\theta+3\theta)^{x+1}} + \frac{a^2(p-1)}{(1+a\theta+\theta)^{a+1}} - \frac{6pa(a+2)}{(1+a\theta+2\theta)^{x+1}} \right. \\
&\quad \left. + \frac{6p}{(1+3a\theta+3\theta)^{x+1}} + \frac{6pa}{(1+2a\theta+2\theta)^{x+1}} - \frac{6p(a+1)^2}{(1+3\theta)^{x+1}} + \frac{6pa(a+1)}{(1+2\theta)^{x+1}} \right. \\
&\quad \left. - \frac{a^2(p-1)}{(1+\theta)^{x+1}} \right)
\end{aligned}$$

Sub-Model: Equation (7) becomes the PMF of the mixed Poisson Weighted Exponential Distribution (Zamani *et al.*, 2014) when the $p = 0$.

Figure 2 shows different shapes of the PCTWED for different parameter values. The figure reveals that the distribution is suitable for unimodal distribution and observations with excess zeros, resembling the shapes of the PDF of the CTWED in Figure 1.

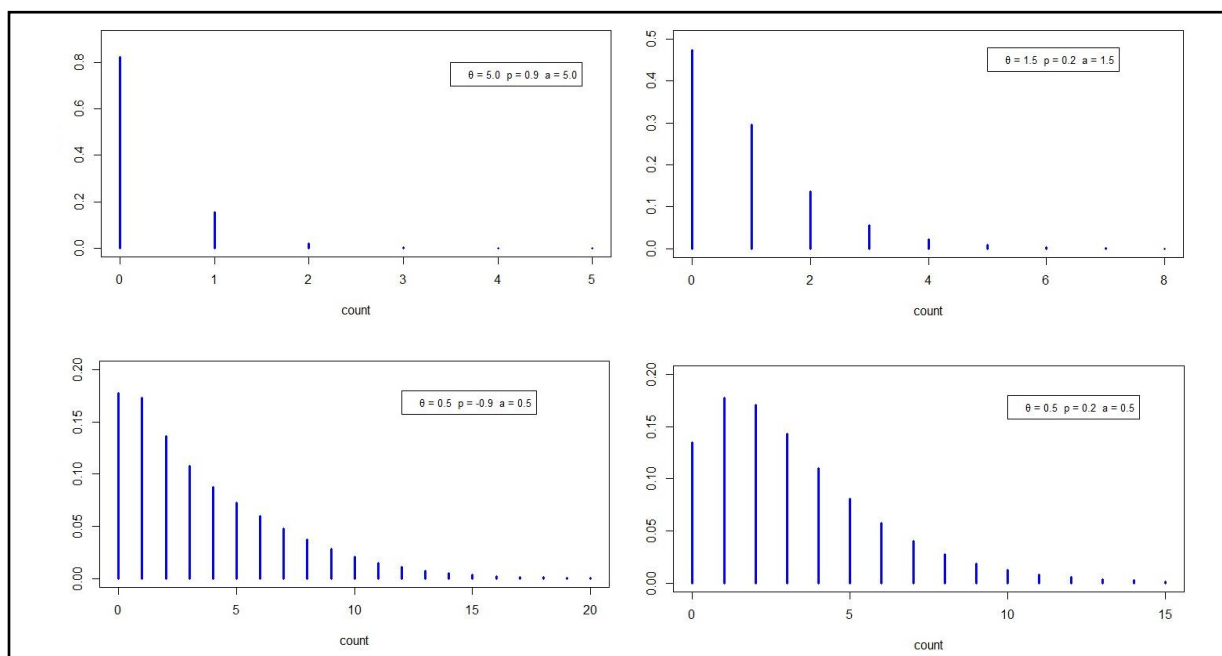


Figure 2. Shapes of the PMF of the PCTWED

The CDF of the PCTWED

Proposition 4: The CDF of a random variable X with the PCTWED is defined in equation (8) as:

$$F(x) = 1 - \left(\frac{6p(a+1)^2}{a^3(1+a\theta+3\theta)^{x+1}} - \frac{6p(a+1)}{a^3(1+2a\theta+3\theta)^{x+1}} + \frac{(p-1)}{a(1+a\theta+\theta)^{x+1}} - \frac{6p(a+1)}{a^2(1+a\theta+2\theta)^{x+1}} + \frac{2p}{a^3(1+3a\theta+3\theta)^{x+1}} + \frac{3p}{a^2(1+2a\theta+2\theta)^{x+1}} - \frac{2p(a+1)^3}{a^3(1+3\theta)^{x+1}} + \frac{3p(a+1)^2}{a^2(1+2\theta)^{x+1}} - \frac{(p-1)(a+1)}{a(1+\theta)^{x+1}} \right) \tag{8}$$

Proof

$$\begin{aligned} F(X) &= P(X \leq x) \\ &= 1 - P(X > x) \\ &= 1 - \sum_{k=x+1}^{\infty} P_k \\ &= 1 - \sum_{k=x+1}^{\infty} \frac{\theta(a+1)}{a^3} \left(\frac{6p(a+1)(a+3)}{(1+a\theta+3\theta)^{k+1}} - \frac{6p(2a+3)}{(1+2a\theta+3\theta)^{k+1}} + \frac{a^2(p-1)}{(1+a\theta+\theta)^{k+1}} - \frac{6pa(a+2)}{(1+a\theta+2\theta)^{k+1}} + \frac{6p}{(1+3a\theta+3\theta)^{k+1}} + \frac{6pa}{(1+2a\theta+2\theta)^{k+1}} - \frac{6p(a+1)^2}{(1+3\theta)^{k+1}} + \frac{6pa(a+1)}{(1+2\theta)^{k+1}} - \frac{a^2(p-1)}{(1+\theta)^{k+1}} \right) \\ &= 1 - \frac{\theta(a+1)}{a^3} \left(\frac{6p(a+1)(a+3)}{\theta(a+3)(1+a\theta+3\theta)^{x+1}} - \frac{6p(2a+3)}{\theta(2a+3)(1+2a\theta+3\theta)^{x+1}} + \frac{a^2(p-1)}{\theta(a+1)(1+a\theta+\theta)^{x+1}} - \frac{6pa(a+2)}{\theta(a+2)(1+a\theta+2\theta)^{x+1}} + \frac{6p}{3\theta(a+1)(1+3a\theta+3\theta)^{x+1}} + \frac{6pa}{2\theta(a+1)(1+2a\theta+2\theta)^{x+1}} - \frac{6p(a+1)^2}{3\theta(1+3\theta)^{x+1}} + \frac{6pa(a+1)}{2\theta(1+2\theta)^{x+1}} - \frac{a^2(p-1)}{2\theta(1+\theta)^{x+1}} \right) \text{Zero-Inflated Poisson} \end{aligned}$$

$$= 1 - \left(\frac{6p(a+1)^2}{a^3(1+a\theta+3\theta)^{x+1}} - \frac{6p(a+1)}{a^3(1+2a\theta+3\theta)^{x+1}} + \frac{(p-1)}{a(1+a\theta+\theta)^{x+1}} - \frac{6p(a+1)}{a^2(1+a\theta+2\theta)^{x+1}} + \frac{2p}{a^3(1+3a\theta+3\theta)^{x+1}} + \frac{3p}{a^2(1+2a\theta+2\theta)^{x+1}} - \frac{2p(a+1)^3}{a^3(1+3\theta)^{x+1}} + \frac{3p(a+1)^2}{a^2(1+2\theta)^{x+1}} - \frac{(p-1)(a+1)}{a(1+\theta)^{x+1}} \right)$$

Mathematical Properties of the PCTWED

Proposition 5: If $f(t)$ is the PDF of the mixing distribution for a discrete random X , the Probability Generating Function (PGF) is obtained as:

$$\begin{aligned} P_x(z) &= \int_0^\infty e^{t(z-1)} f(t) dt \\ &= \int_0^\infty e^{t(z-1)} \frac{\theta(a+1)}{a^3} [6p(a+1)(a+3)e^{-(a+3)\theta t} - 6p(2a+3)e^{-(2a+3)\theta t} \\ &\quad + a^2(p-1)e^{-(a+1)\theta t} - 6pa(a+2)e^{-(a+2)\theta t} + 6pe^{-3(a+1)\theta t} + 6pae^{-2(a+1)\theta t} \\ &\quad - 6p(a+1)^2 e^{-3\theta t} + 6pa(a+1)e^{-2\theta t} - a^2(p-1)e^{-\theta t}] dt \\ &= \frac{\theta(a+1)}{a^3} \int_0^\infty (6p(a+1)(a+3)e^{-(1+a\theta+3\theta-z)t} - 6p(2a+3)e^{-(1+2a\theta+3\theta-z)t} \\ &\quad + a^2(p-1)e^{-(1+a\theta+\theta-z)t} - 6pa(a+2)e^{-(1+a\theta+2\theta-z)t} + 6pe^{-(1+3a\theta+3\theta-z)t} \\ &\quad + 6pae^{-(1+2a\theta+\theta-z)t} - 6p(a+1)^2 e^{-(1+3\theta-z)t} + 6pa(a+1)e^{-(1+2\theta-z)t} \\ &\quad - a^2(p-1)e^{-(1+\theta-z)t}) dt \end{aligned}$$

Therefore, the PGF is obtained in equation (9) as

$$\begin{aligned} P_x(z) &= \frac{\theta(a+1)}{a^3} \left(\frac{6p(a+1)(a+3)}{(1+a\theta+3\theta-z)} - \frac{6p(2a+3)}{(1+2a\theta+3\theta-z)} + \frac{a^2(p-1)}{(1+a\theta+\theta-z)} - \frac{6pa(a+2)}{(1+a\theta+2\theta-z)} + \frac{6p}{(1+3a\theta+3\theta-z)} + \frac{6pa}{(1+2a\theta+\theta-z)} - \frac{6p(a+1)^2}{(1+3\theta-z)} + \frac{6pa(a+1)}{(1+2\theta-z)} - \frac{a^2(p-1)}{(1+\theta-z)} \right) \end{aligned} \tag{9}$$

The Moment Generating Function (MGF) is obtained in equation (10) by replacing z with e^z in equation (9).

$$\begin{aligned} M_X(z) &= \frac{\theta(a+1)}{a^3} \left(\frac{6p(a+1)(a+3)}{(1+a\theta+3\theta-e^z)} - \frac{6p(2a+3)}{(1+2a\theta+3\theta-e^z)} + \frac{a^2(p-1)}{(1+a\theta+\theta-e^z)} - \frac{6pa(a+2)}{(1+a\theta+2\theta-e^z)} + \frac{6p}{(1+3a\theta+3\theta-e^z)} + \frac{6pa}{(1+2a\theta+\theta-e^z)} - \frac{6p(a+1)^2}{(1+3\theta-e^z)} + \frac{6pa(a+1)}{(1+2\theta-e^z)} - \frac{a^2(p-1)}{(1+\theta-e^z)} \right) \end{aligned} \tag{10}$$

The first four central moments are obtained in equations (11) - (14) as:

$$E(X) = \frac{(12-2p)a^4 + (102-15p)a^3 + (318-34p)a^2 + (432-38p)a + 216-19p}{6\theta(a+1)(a+2)(a+3)(2a+3)} \tag{11}$$

$$E(X^2) = \left(\frac{\theta(1+a)}{a^3}\right) \left(\frac{12p(a+1)(a+3)}{(a\theta+3\theta)^3} + \frac{6p(a+1)(a+3)}{(a\theta+3\theta)^2} - \frac{12p(2a+3)}{(2a\theta+3\theta)^3} - \frac{6p(2a+3)}{(2a\theta+3\theta)^2} + \frac{2a^2(p-1)}{(a\theta+\theta)^3} + \frac{a^2(p-1)}{(a\theta+\theta)^2} - \frac{12pa(a+2)}{(a\theta+2\theta)^3} - \frac{6pa(a+2)}{(a\theta+2\theta)^2} + \frac{12p}{(3a\theta+3\theta)^3} + \frac{6p}{(3a\theta+3\theta)^2} + \frac{12pa}{(2a\theta+2\theta)^3} + \frac{6pa}{(2a\theta+2\theta)^2} - \frac{8p(1+a)^2-27pa(1+a)+36a^2(p-1)}{18\theta^3} - \frac{4p(1+a)^2-9pa(1+a)+6a^2(p-1)}{6\theta^2}\right) \tag{12}$$

$$E(X^3) = \left(\frac{\theta(a+1)}{a^3}\right) \left(\frac{36p}{(3a\theta+3\theta)^4} + \frac{36p}{(3a\theta+3\theta)^3} + \frac{6p}{(3a\theta+3\theta)^2} - \frac{36p(2a+3)}{(2a\theta+3\theta)^4} - \frac{36p(2a+3)}{(2a\theta+3\theta)^3} - \frac{6p(2a+3)}{(2a\theta+3\theta)^2} + \frac{36pa}{(2a\theta+2\theta)^4} + \frac{36pa}{(2a\theta+2\theta)^3} + \frac{6pa}{(2a\theta+2\theta)^2} + \frac{36p(a+1)(a+3)}{(a\theta+3\theta)^4} + \frac{36p(a+1)(a+3)}{(a\theta+3\theta)^3} + \frac{6p(a+1)(a+3)}{(a\theta+3\theta)^2} - \frac{36pa(a+2)}{(a\theta+2\theta)^4} - \frac{36pa(a+2)}{(a\theta+2\theta)^3} - \frac{6pa(a+2)}{(a\theta+2\theta)^2} + \frac{6a^2(p-1)}{(a\theta+\theta)^4} + \frac{6a^2(p-1)}{(a\theta+\theta)^3} + \frac{a^2(p-1)}{(a\theta+\theta)^2} - \frac{216a^2(p-1)-81pa(a+1)+16p(a+1)^2}{36\theta^4} - \frac{36a^2(p-1)-27pa(a+1)+8p(a+1)^2}{6\theta^3} - \frac{4p(a+1)^2-9pa(a+1)+6a^2(p-1)}{6\theta^2}\right) \tag{13}$$

$$E(X^4) = \left(\frac{\theta(a+1)}{a^3}\right) \left(\frac{144p}{(3a\theta+3\theta)^5} + \frac{216p}{(3a\theta+3\theta)^4} + \frac{84p}{(3a\theta+3\theta)^3} + \frac{6p}{(3a\theta+3\theta)^2} - \frac{144p(2a+3)}{(2a\theta+3\theta)^5} - \frac{216p(2a+3)}{(2a\theta+3\theta)^4} - \frac{84p(2a+3)}{(2a\theta+3\theta)^3} - \frac{6p(2a+3)}{(2a\theta+3\theta)^2} + \frac{144pa}{(2a\theta+2\theta)^5} + \frac{216pa}{(2a\theta+2\theta)^4} + \frac{84pa}{(2a\theta+2\theta)^3} + \frac{6pa}{(2a\theta+2\theta)^2} + \frac{144p(a+1)(a+3)}{(a\theta+3\theta)^5} + \frac{216p(a+1)(a+3)}{(a\theta+3\theta)^4} + \frac{84p(a+1)(a+3)}{(a\theta+3\theta)^3} + \frac{6p(a+1)(a+3)}{(a\theta+3\theta)^2} - \frac{144pa(a+2)}{(a\theta+2\theta)^5} - \frac{216pa(a+2)}{(a\theta+2\theta)^4} - \frac{84pa(a+2)}{(a\theta+2\theta)^3} - \frac{6pa(a+2)}{(a\theta+2\theta)^2} + \frac{24a^2(p-1)}{(a\theta+\theta)^5} + \frac{36a^2(p-1)}{(a\theta+\theta)^4} + \frac{14a^2(p-1)}{(a\theta+\theta)^3} + \frac{a^2(p-1)}{(a\theta+\theta)^2} - \frac{32p(a+1)^2+1296a^2(p-1)-243pa(a+1)}{54\theta^5} - \frac{16p(a+1)^2+216a^2(p-1)-81pa(a+1)}{6\theta^4} - \frac{56p(a+1)^2+252a^2(p-1)-189pa(a+1)}{18\theta^3} - \frac{4p(a+1)^2+6a^2(a-1)-9pa(a+1)}{6\theta^2}\right) \tag{14}$$

Skewness and Kurtosis

The skewness and kurtosis for the PCTWED are obtained from the central moments (De Jong & Heller, 2008) as:

$$S_k = \frac{E(X^3)-3E(X^2)E(X)+2(E(X))^3}{(Var(X))^{\frac{3}{2}}}$$

$$Kurt = \frac{E(X^4)-4E(X^3)E(X)+6E(X^2)(E(X))^2-3(E(X))^4}{(Var(X))^2}$$

Tables 1 – 3 show simulated Skewness, Kurtosis, and Dispersion Index for some parameters distribution.

Table 1. Skewness for some parameters of the PCTWED

	a = 0.5			a = 2.5			a = 7.5		
	θ = 0.1	θ = 2.0	θ = 10.0	θ = 0.1	θ = 2.0	θ = 10.0	θ = 0.1	θ = 2.0	θ = 10.0
p = -0.9	1.68	2.44	4.06	1.76	3.08	5.82	1.94	4.27	8.80
p = -0.5	1.82	2.48	4.15	1.89	3.13	5.97	2.03	4.36	9.04
p = 0.0	2.00	2.50	4.25	2.02	3.18	6.17	2.10	4.48	9.38
p = 0.5	2.10	2.46	4.35	2.05	3.20	6.38	2.06	4.59	9.76
p = 0.9	1.77	2.34	4.41	1.74	3.18	6.55	1.80	4.67	10.09

Table 2. Kurtosis for some parameters of the PCTWED

	a = 0.5			a = 2.5			a = 7.5		
	θ = 0.1	θ = 2.0	θ = 10.0	θ = 0.1	θ = 2.0	θ = 10.0	θ = 0.1	θ = 2.0	θ = 10.0
p = -0.9	6.60	10.61	21.61	7.06	14.46	39.12	7.96	23.38	82.65
p = -0.5	7.56	10.98	22.30	7.90	14.83	40.80	8.62	24.15	86.98
p = 0.0	9.02	11.25	23.06	9.09	15.13	43.03	9.39	25.06	93.01
p = 0.5	10.40	10.97	23.59	9.93	15.01	45.33	9.59	25.76	99.84
p = 0.9	8.72	9.82	23.68	8.17	14.34	47.16	7.96	26.02	105.97

Table 3. Dispersion Index for some parameters of the PCTWED

	a = 0.5			a = 2.5			a = 7.5		
	θ = 0.1	θ = 2.0	θ = 10.0	θ = 0.1	θ = 2.0	θ = 10.0	θ = 0.1	θ = 2.0	θ = 10.0
p = -0.9	9.86	1.44	1.09	4.80	1.19	1.04	2.56	1.08	1.02
p = -0.5	8.99	1.40	1.08	4.42	1.17	1.03	2.41	1.07	1.01
p = 0.0	7.67	1.33	1.07	3.86	1.14	1.03	2.18	1.06	1.01
p = 0.5	6.00	1.25	1.05	3.14	1.11	1.02	1.88	1.04	1.01
p = 0.9	4.35	1.17	1.03	2.44	1.07	1.01	1.59	1.03	1.01

Remarks:

For fixed p and a , skewness and kurtosis increase while the dispersion index decreases as θ increases. When p and θ are fixed, skewness and kurtosis increase while the dispersion index reduces as a increases. For fixed a and θ , skewness, kurtosis increase while the dispersion index decreases as p increases in most cases. Both skewness and kurtosis peak at $p = 0.5$.

Maximum Likelihood Estimation of the PCTWED

Given random samples of size n drawn from the PCTWED with (θ, p, a) as defined in equation (7), the log-likelihood function for the distribution is obtained in equation (15) as:

$$\ell = \sum_{i=1}^n \log \left(\frac{\theta(a+1)}{a^3} \left(\frac{6p(a+1)(a+3)}{(1+a\theta+3\theta)^{x+1}} - \frac{6p(2a+3)}{(1+2a\theta+3\theta)^{x+1}} + \frac{a^2(p-1)}{(1+a\theta+\theta)^{x+1}} - \frac{6pa(a+2)}{(1+a\theta+2\theta)^{x+1}} + \frac{6p}{(1+3a\theta+3\theta)^{x+1}} + \frac{6pa}{(1+2a\theta+2\theta)^{x+1}} - \frac{6p(a+1)^2}{(1+3\theta)^{x+1}} + \frac{6pa(a+1)}{(1+2\theta)^{x+1}} - \frac{a^2(p-1)}{(1+\theta)^{x+1}} \right) \right) \tag{15}$$

Estimators for (θ, p, a) denoted with $(\hat{\theta}, \hat{p}, \hat{a})$ are the solutions for the log-likelihood function. The equations form a non-linear equation system that can only be solved numerically. This research uses the optimr function (Nash *et al.*, 2019) in the **R language** (R-Core Team, 2020) is used.

Zero-Inflated PCTWED

Proposition 6. If a discrete random variable X has a PCTWED with PMF P_x and its realization at $x = 0$ as P_0 . If the zero-inflation parameter is denoted with π , then a discrete random variable X_Z has a zero-inflated PCTWED if its PMF is defined in equation (16) as:

$$P_x^Z = \begin{cases} \pi + (1 - \pi)P_0, & x = 0 \\ (1 - \pi)P_x, & x = 1, 2, 3, \dots \end{cases} \tag{16}$$

where

$$P_x = \frac{\theta(a+1)}{a^3} \left(\frac{6p(a+1)(a+3)}{(1+a\theta+3\theta)^{x+1}} - \frac{6p(2a+3)}{(1+2a\theta+3\theta)^{x+1}} + \frac{a^2(p-1)}{(1+a\theta+\theta)^{x+1}} - \frac{6pa(a+2)}{(1+a\theta+2\theta)^{x+1}} + \frac{6p}{(1+3a\theta+3\theta)^{x+1}} + \frac{6pa}{(1+2a\theta+2\theta)^{x+1}} - \frac{6p(a+1)^2}{(1+3\theta)^{x+1}} + \frac{6pa(a+1)}{(1+2\theta)^{x+1}} - \frac{a^2(p-1)}{(1+\theta)^{x+1}} \right)$$

and

$$P_0 = \frac{\theta(a+1)}{a^3} \left(\frac{6p(a+1)(a+3)}{(1+a\theta+3\theta)} - \frac{6p(2a+3)}{(1+2a\theta+3\theta)} + \frac{a^2(p-1)}{(1+a\theta+\theta)} - \frac{6pa(a+2)}{(1+a\theta+2\theta)} + \frac{6p}{(1+3a\theta+3\theta)} + \frac{6pa}{(1+2a\theta+2\theta)} - \frac{6p(a+1)^2}{(1+3\theta)} + \frac{6pa(a+1)}{(1+2\theta)} - \frac{a^2(p-1)}{(1+\theta)} \right)$$

Proof

The proof is obtained by appropriate substitution. Hence the result.

Mathematical Properties of the Zero-Inflated PCTWED

If the PGF of the PCTWED is denoted with $P_x(z)$, then the PGF of the Zero-Inflated PCTWED denoted by $P_x^Z(z)$ is obtained using: $P_x^Z(z) = (1 - \pi)P_x(z)$

Hence, the PGF of the Zero-Inflated PCTWED is given in equation (17) as:

$$P_x^Z(z) = (1 - \pi) \left(\frac{\theta(a+1)}{a^3} \left(\frac{6p(a+1)(a+3)}{(1+a\theta+3\theta-z)} - \frac{6p(2a+3)}{(1+2a\theta+3\theta-z)} + \frac{a^2(p-1)}{(1+a\theta+\theta-z)} - \frac{6pa(a+2)}{(1+a\theta+2\theta-z)} + \frac{6p}{(1+3a\theta+3\theta-z)} + \frac{6pa}{(1+2a\theta+\theta-z)} - \frac{6p(a+1)^2}{(1+3\theta-z)} + \frac{6pa(a+1)}{(1+2\theta-z)} - \frac{a^2(p-1)}{(1+\theta-z)} \right) \right) \tag{17}$$

The corresponding MGF is therefore expressed in equation (18) as:

$$M_x^Z(z) = (1 - \pi) \left(\frac{\theta(a+1)}{a^3} \left(\frac{6p(a+1)(a+3)}{(1+a\theta+3\theta-e^t)} - \frac{6p(2a+3)}{(1+2a\theta+3\theta-e^t)} + \frac{a^2(p-1)}{(1+a\theta+\theta-e^t)} - \frac{6pa(a+2)}{(1+a\theta+2\theta-e^t)} + \frac{6p}{(1+3a\theta+3\theta-e^t)} + \frac{6pa}{(1+2a\theta+\theta-e^t)} - \frac{6p(a+1)^2}{(1+3\theta-e^t)} + \frac{6pa(a+1)}{(1+2\theta-z)} - \frac{a^2(p-1)}{(1+\theta-e^t)} \right) \right) \tag{18}$$

If the r^{th} raw moment of the PCTWED is denoted by $E(X^r)$, then the r^{th} raw moment of the zero-inflated PCTWED is generally defined as:

$$m_r = E(X_r^r) = (1 - \pi)E(X^r)$$

Given the first four raw moments of the PCTWED as $E(X)$, $E(X^2)$, $E(X^3)$, and $E(X^4)$ as obtained in equations (11) to (14), the first four raw moments of the zero-inflated PCTWED are given in equations (19) to (22).

$$m_1 = (1 - \pi)E(X) \tag{19}$$

$$m_2 = (1 - \pi)E(X^2) \tag{20}$$

$$m_3 = (1 - \pi)E(X^3) \tag{21}$$

$$m_4 = (1 - \pi)E(X^4) \tag{22}$$

The variance, dispersion index, skewness, and kurtosis are obtained respectively from equations (19) to (22) as:

$$\begin{aligned} \text{Var}(X_Z) &= m_2 - [m_1]^2 \\ DI &= \frac{\text{Var}(X_Z)}{m_1} \\ S_k &= \frac{m_3 - 3m_2m_1 + 2(m_1)^3}{(\text{Var}(X_Z))^{\frac{3}{2}}} \\ kurt &= \frac{m_4 - 4m_3m_1 + 6m_2(m_1)^2 - 3(m_1)^4}{(\text{Var}(X_Z))^2} \end{aligned}$$

Maximum Likelihood Estimation of the Parameter of Zero-Inflated PCTWED

If a random variable X is assumed to follow the Zero-Inflated PCTWED with PMF P_x^Z indexed with parameter (θ, p, a, π) , then the distribution parameters can be estimated using the MLE method. The likelihood function is defined as:

$$\mathcal{L}(\theta, p, a, \pi) = \prod_{n_0} (\pi + (1 - \pi)P_0) \prod_{n_1} ((1 - \pi)P(x > 0))$$

where: n_0 is the frequency of zero in the observation; n_1 is the frequency of non-zero observations; the sample size $n = (n_0 + n_1)$, P_0 is the realization of P_x at $x = 0$. The log-likelihood function is obtained as:

$$\begin{aligned} \ell &= n_0 \ln(\pi + (1 - \pi)P_0) + n_1 \ln(1 - \pi) + \left(\sum_{n_1} \ln(P_x) \right) \\ &= n_0 \ln \left(\pi + (1 - \pi) \left(\frac{\theta(a+1)}{a^3} \left(\frac{6p(a+1)(a+3)}{(1+a\theta+3\theta)} - \frac{6p(2a+3)}{(1+2a\theta+3\theta)} + \frac{a^2(p-1)}{(1+a\theta+\theta)} - \frac{6pa(a+2)}{(1+a\theta+2\theta)} + \frac{6p}{(1+3a\theta+3\theta)} + \right. \right. \right. \\ &\quad \left. \left. \frac{6pa}{(1+2a\theta+2\theta)} - \frac{6p(a+1)^2}{(1+3\theta)} + \frac{6pa(a+1)}{(1+2\theta)} - \frac{a^2(p-1)}{(1+\theta)} \right) \right) + n_1 \ln(1 - \pi) + \sum_{n_1} \ln \left(\frac{\theta(a+1)}{a^3} \left(\frac{6p(a+1)(a+3)}{(1+a\theta+3\theta)^{x+1}} - \right. \right. \\ &\quad \left. \left. \frac{6p(2a+3)}{(1+2a\theta+3\theta)^{x+1}} + \frac{a^2(p-1)}{(1+a\theta+\theta)^{x+1}} - \frac{6pa(a+2)}{(1+a\theta+2\theta)^{x+1}} + \frac{6p}{(1+3a\theta+3\theta)^{x+1}} + \frac{6pa}{(1+2a\theta+2\theta)^{x+1}} - \frac{6p(a+1)^2}{(1+3\theta)^{x+1}} + \frac{6pa(a+1)}{(1+2\theta)^{x+1}} - \right. \right. \\ &\quad \left. \left. \frac{a^2(p-1)}{(1+\theta)^{x+1}} \right) \right) \\ \frac{\partial \ell}{\partial \pi} &= \frac{n_0(1-P_0)}{\pi+(1-\pi)P_0} - \frac{n_1}{(1-\pi)} \\ \hat{\pi} &= \frac{n_0}{n_1} - \frac{n_1}{n} \left(\frac{P_0}{1-P_0} \right) \end{aligned}$$

The MLE for parameters space (θ, p, a, π) are obtained numerically by solving $\frac{\partial \ell}{\partial \pi} = 0$, $\frac{\partial \ell}{\partial \theta} = 0$, $\frac{\partial \ell}{\partial p} = 0$, and $\frac{\partial \ell}{\partial a} = 0$. This is done using different algorithms that come with the **optimr** package (Nash *et al.*, 2019) in **R language** (R-Core Team, 2020).

APPLICATIONS

Data

Five count datasets characterized by many zeros are considered to assess the performance of the PCTWED and zero-inflated PCTWED. The new propositions are compared with the Poisson and negative binomial distributions (along with their respective zero-inflated forms). Dataset I consists of the number of claims on motorcycle insurance from WASA (a Swedish Insurance Firm) from 1994 – 1998. The data has been previously utilized (Omari *et al.*, 2018) on count distributions. The second dataset is the number of automobile insurance policies in Australia between 2004 and 2005, as previously presented in De Jong and Heller (2008). Dataset III comprises of frequency of claim insurance in a Belgium firm in 1993. The data was used to assess claim distributions (Zamani *et al.*, 2014). The fourth dataset is the frequency of claims of 10, 814 policyholders of a Turkish insurance firm between 2012 and 2014. The data were assessed on Poisson-related distributions (Meytrianti *et al.*, 2019). The fifth dataset is the yeast cell counts per square, as previously examined on Poison Lindley distribution (Shanker & Hagos, 2015). All five datasets are dispersed and positively skewed, with very high percentages of zero counts (table 4).

Table 4. Summary of Datasets

X	Dataset I	Dataset II	Dataset III	Dataset IV	Dataset V
0	63878	63232	57178	8544	128
1	643	4333	5617	1796	37
2	27	271	446	370	18
3		18	50	81	3
4		2	8	23	1
% of 0	98.96	93.19	90.33	79.01	68.45
Dispersion Index	1.07	1.08	1.08	1.26	1.32
Kurtosis	118.26	18.50	14.59	7.71	2.80
Skewness	10.46	4.07	3.52	2.56	1.75

RESULTS AND DISCUSSION

With the least values of the chi-square statistic and $-LL$, the proposed distribution (PCTWED) provide the best fit to the first dataset (table 5) while its zero-inflated form (ZI-PCTWED) performs worst. It is also observed that the negative binomial distribution (another mixed Poisson distribution) performs better than its zero-inflated form, while the ZIP performs better than the Poisson distribution.

For the second dataset (table 6), the PCTWED also gives the best fit with the lowest values of both $-LL$ and chi-square statistic. Like the first data, the ZiNB also performs better than the classical negative binomial distribution, while the ZIP performs better than the Poisson distribution.

Table 7 shows that the ZiNB best fits the third dataset; the PCTWED follows this. The ZI-PCTWED gives the worst fit to the data with the highest values of both $-LL$ and chi-square.

Tables 8 and 9 show that the new proposition (PCTWED) best fits both datasets IV and V. The negative binomial provides a better fit than its zero-inflated form, while the ZIP gives a better fit than

the Poisson distribution for both datasets. Dataset V has the lowest percentage of zero counts among the five datasets assessed. The ZI-PCTWED gives a relatively better fit in this dataset than the negative binomial distribution and its zero-inflated form.

Table 5. Data I (Swedish Claim Frequency)

X	Freq.	PCTWED	ZI-PCTWED	Poisson	ZIP	Neg. Bin.	ZiNB
0	63878	63878.13	63877.74	63854.63	63878.11	63877.73	63893.56
1	643	642.83	632.94	689.63	643.61	644.62	629.97
2	27	26.76	37.38	3.72	25.58	24.43	23.26
<i>Estimates</i>	θ	89.097	504.897	0.011	0.080	0.154	0.003
	p	-4.759	-0.761		0.864	0.934	0.926
	a	-4.926	-3.590				-40.710
	π		-54.232				
	$-LL$	3840.44	3853.79	3872.00	3840.88	3841.50	3857.32
	<i>Chi-Square</i>	0.00	3.04	148.64	0.08	0.27	0.88

Table 6. Data II (Australian Claim Frequency)

X	Freq.	PCTWED	ZI-PCTWED	Poisson	ZIP	Neg. Bin.	ZiNB
0	63232	63230.90	63239.55	63091.61	63230.49	63230.60	63317.89
1	4333	4332.84	4310.28	4593.07	4325.83	4330.57	4252.49
2	271	273.98	305.84	167.19	286.59	276.48	261.98
3	18	17.14	0.33	4.06	12.66	17.22	21.45
4	2	1.07	0.00	0.07	0.42	1.06	1.97
<i>Estimates</i>	θ	15.167	387.325	0.073	0.133	1.157	0.007
	p	-0.045	-5.445		0.451	0.941	0.878
	a	9.406	-4.253				-76.060
	π		-135.921				
	$-LL$	18049.64	18148.72	18101.50	18052.20	18049.68	18105.58
	<i>Chi-Square</i>	0.89	2585.95	177.66	9.07	0.98	48.15

Table 7. Data III (Belgium Claim Frequency)

X	Freq.	PCTWED	ZI-PCTWED	Poisson	ZIP	Neg. Bin.	ZiNB
0	57178	57178.64	57187.37	56949.763	57177.48	57188.34	57249.63
1	5617	5598.70	5587.71	6019.590	5584.80	5581.31	5558.90
2	446	477.08	523.00	318.135	504.87	485.28	438.37
3	50	40.67	0.91	11.209	30.43	40.47	45.91
4	8	3.57	0.01	0.296	1.38	3.30	5.40
<i>Estimates</i>	θ	9.573	271.157	0.106	0.181	1.279	0.008
	p	0.354	-6.343		0.415	0.924	0.843
	a	13.223	-4.312				-71.130
	π		-111.862				
	$-LL$	22063.30	22311.68	22150.54	22075.30	22064.31	22136.57
	<i>Chi-Square</i>	9.74	10122.64	413.84	51.55	12.33	2.44

Table 8. Data IV (Turkish Claim Frequency)

X	Freq.	PCTWED	ZI-PCTWED	Poisson	ZIP	Neg. Bin.	ZiNB
0	8544	8545.14	8544.04	8292.42	8544.19	8543.47	8561.78
1	1796	1789.69	1771.07	2201.64	1759.23	1795.62	1807.66
2	370	380.67	492.20	292.27	430.75	375.71	331.89
3	81	78.92	6.34	25.87	70.31	78.50	81.03
4	23	15.78	0.33	1.72	8.61	16.39	22.23
<i>Estimates</i>	θ	4.353	41.119	0.266	0.490	1.009	0.006
	p	-0.506	-4.526		0.458	0.792	0.635
	a	12.831	-4.408				-82.430
	π		-19.730				
	$-LL$	7029.56	7334.26	7153.16	7038.91	7029.71	7057.06
	<i>Chi-Square</i>	2.68	2452.35	484.41	35.02	2.84	4.52

Table 9. Data V (Yeast cell counts per square)

X	Freq.	PCTWED	ZI-PCTWED	Poisson	ZIP	Neg. Bin.	ZiNB
0	128	127.31	127.48	118.06	128.00	126.73	180.68
1	37	38.36	35.34	54.30	38.35	42.08	4.51
2	18	16.72	21.26	12.49	15.49	12.84	1.15
3	3	3.86	2.51	1.91	4.17	3.80	0.39
4	1	0.65	0.35	0.22	0.84	1.11	0.15
<i>Estimates</i>	θ	9.003	97.851	0.4609	0.808	1.195	0.004
	p	-12.457	-0.232		0.431	0.722	0.490
	a	0.013	-1.467				-11.720
	π		-31.216				
	$-LL$	168.60	170.31	173.83	168.80	170.02	172.05
	<i>Chi-Square</i>	0.53	1.92	12.16	0.81	2.88	517.31

CONCLUSION

The skewness, kurtosis, and dispersion index for some parameter combinations are presented to assess the behaviour of the distributions. The shapes of the PCTWED show that it can be effectively utilized to model observations with an unusual frequency of zeros.

Performances of the new propositions are assessed on five count observations with varying percentages of zero counts. Comparisons are made with the Poisson and negative binomial distributions (along with the irrespective zero-inflated forms). The maximum likelihood estimation using different algorithms that come with the optimr package in R-language is used to provide estimates for the parameters of the distributions. The chi-square goodness of fits and the $-LL$ are used for model selection. Results show that the PCTWED outperforms its zero-inflated form in particular and all the competing

distributions in most cases. The classical negative binomial distribution provides a better fit for datasets with above 70% of zero counts than its zero-inflated form. In contrast, the zero-inflated Poisson outperforms the Poisson distribution in all cases.

The finding shows that mixed Poisson distributions and the negative binomial distribution (when the gamma distribution is used as the mixing distribution) are naturally suitable to model observation with higher than usual zero counts in count observation. Therefore, it is unnecessary to obtain their zero-inflated form to model observations with many zero counts.

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